

UNIVERSITY OF MASSACHUSETTS  
DEPARTMENT OF MATHEMATICS AND STATISTICS  
ADVANCED EXAM - STATISTICS (II)  
Tuesday, January 25, 2021

Work all problems and show all work. Explain your answers. State the theorems used whenever possible. 70 points are required to pass.

1. Let  $\{Y_n\}_{n \geq 1}$  be a sequence of real-valued random variables and  $Y$  be another real-valued random variable. Suppose that  $Y_n$  has distribution function  $F_n(y)$  for each  $n$  and  $Y$  has distribution function  $F(y)$ .
  - (a) (3 points) State the definition of convergence in probability (denoted as  $Y_n \xrightarrow{P} Y$ ).
  - (b) (3 points) State the definition of convergence in quadratic mean (denoted as  $Y_n \xrightarrow{qm} Y$ ).
  - (c) (3 points) State the definition of convergence in distribution (denoted as  $Y_n \xrightarrow{d} Y$ ).
  - (d) (5 points) Show that  $Y_n \xrightarrow{qm} c$  if and only if  $EY_n \rightarrow c$  and  $VarY_n \rightarrow 0$  for a constant  $c$ .
  - (e) (6 points) Show that if  $Y_n \xrightarrow{d} Y$ , then  $Y_n = O_p(1)$  (i.e., for every  $\epsilon > 0$ , there exist  $M$  and  $N$  such that  $P(|Y_n| > M) < \epsilon$  for  $n > M$ ).
2. Suppose that  $X_1, \dots, X_n$  are independent but not identically distributed random variables with  $E(X_i) = \mu$  and  $Var(X_i) = \sigma_i^2$  where  $i = 1, \dots, n$ . Consider two estimators for  $\mu$  :

$$\hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n X_i \text{ and } \hat{\mu}_2 = \frac{\sum_{i=1}^n X_i / \sigma_i^2}{\sum_{j=1}^n 1 / \sigma_j^2}.$$

- (a) (5 points) What is a sufficient condition for which  $\hat{\mu}_1$  is consistent for  $\mu$ ?
  - (b) (5 points) What is a sufficient condition for which  $\hat{\mu}_2$  is consistent for  $\mu$ ?
  - (c) (10 points) Which estimator is preferable for estimating  $\mu$ ? Justify your answer.
3. Answer the following questions.
    - (a) (5 points) Suppose  $X_1, X_2, \dots$ , is a stationary sequence with  $E(X_i) = \mu$  and  $Var(X_i) = \sigma^2 < \infty$ . Then we know that

$$\sqrt{n}(\bar{X}_n - \mu) \xrightarrow{d} N(0, \tau^2),$$

where  $\tau^2$  is finite and

$$\tau^2 = \lim_{n \rightarrow \infty} \left[ \sigma^2 + \frac{2}{n} \sum_{i=1}^n (n-i) Cov(X_1, X_{1+i}) \right].$$

Suppose that we further assume that  $X_1, X_2, \dots$ , is a stationary  $m$ -dependent sequence with  $E(X_i) = \mu$  and  $Var(X_i) = \sigma^2 < \infty$ . Then what is the limiting distribution of  $\sqrt{n}(\bar{X}_n - \mu)$ ?

- (b) (10 points) Suppose  $X_0, X_1, \dots$  is an iid sequence of Bernoulli trials with success probability  $p$ . Suppose  $X_i$  is the indicator of your team's success on rally  $i$  in a volleyball game. (Note: This is a completely unrealistic model, since the serving team is always at a disadvantage when evenly matched teams play.) Your team scores a point each time it has a success that follows another success. Let  $S_n = \sum_{i=1}^n X_{i-1}X_i$  denote the number of points your team scores by time  $n$ . Find the asymptotic distribution of  $S_n$ .
4. Let  $Y_1, \dots, Y_n$  be independent and identically distributed from a distribution with density function  $f_\theta(y) = \theta/y^{\theta+1}$  where  $y > 1$  and  $\theta > 2$ . Note that  $E(Y_i) = \theta/(\theta - 1)$  and  $Var(Y_i) = \theta/[(\theta - 2)(\theta - 1)^2]$ . Assume that the regularity conditions necessary for asymptotic consistency and efficiency are satisfied.
- (a) (5 points) Obtain the maximum likelihood estimator of  $\theta$ , denoted as  $\hat{\theta}_n$ , and find the limiting distribution of  $\hat{\theta}_n$ .
- (b) (9 points) Obtain the Bayesian estimator  $\hat{\theta}_n^B$  equal to the posterior mean of  $\theta$  under the Jeffreys prior, and find the limiting distribution of  $\hat{\theta}_n^B$ .
- (c) (9 points) Consider an estimator  $\tilde{\theta}_n = (n + c) / \sum_{i=1}^n \log Y_i$  where  $c$  is a constant. Find the limiting distribution of  $\tilde{\theta}_n$ . [Hint]  $E(\log Y_i) = 1/\theta$  and  $Var(\log Y_i) = 1/\theta^2$ .
5. Suppose  $Y_n \sim \text{Binomial}(n, p)$ . We want to test  $H_0 : p = p_0$  against  $H_1 : p > p_0$ .

- (a) (3 points) Prove that under  $H_0 : p = p_0$ ,

$$\frac{\sqrt{n}(Y_n/n - p_0)}{\sqrt{p_0(1 - p_0)}} \xrightarrow{d} Z,$$

where  $Z$  is the standard normal distribution with mean 0 and variance 1.

- (b) (3 points) Suppose that we reject  $H_0 : p = p_0$  whenever

$$Y_n/n \geq p_0 + u_\alpha \sqrt{p_0(1 - p_0)/n}$$

where  $u_\alpha$  is the  $1 - \alpha$  quantile of the standard normal distribution. Show that this test has asymptotic level  $\alpha$  using (a).

- (c) (10 points) Find a sufficient condition under which the following asymptotic result holds for the alternatives  $\{p_n\}$  satisfying  $p_n > p_0$  for all  $n$ :

$$\frac{\sqrt{n}(Y_n/n - p_n)}{\sqrt{p_n(1 - p_n)}} \xrightarrow{d} Z,$$

where  $p_n$  means that the success probability depends on  $n$ .

- (d) (6 points) Suppose that  $p_n > p_0$  for all  $n$  and  $\sqrt{n}(p_n - p_0) \rightarrow \delta > 0$ . Let  $Power_n(p_n)$  be the power of this test against the alternative  $p_n$ . Show that

$$Power_n(p_n) \rightarrow \Phi \left( \frac{\delta}{\sqrt{p_0(1 - p_0)}} - u_\alpha \right) \text{ as } n \rightarrow \infty$$

where  $\Phi(x)$  denotes the cumulative distribution function of the standard normal distribution.