- 1. **1a**; **problem 1.22** Below is the summary information (edited) from the regression (using R output); code at end of solution as is code and output for SAS.
  - a) The estimated regression function is E(Y) = 168.60000 + 2.03438 \* X, where Y is hardness and X is time. The plot of data and fitted line is below. There are two ways to interpret the question of whether the linear regression supplies a good fit. One is whether a straight line does a good job of modeling the expected value. The answer to that seems to be yes from graphical inspection. A second way, and different way, to view the question is how "tight" the fit is around the line. There is clearly variability in hardness at a fixed time. The question of how this variability relates to how good the fit is will depend on how the fit will be used.
  - b) This is the fitted value at X = 40 is  $b_0 + b_1 40$ , which equals 249.975. You can get this directly or use the predicted value in the output for a case with X = 40.
  - c) This is just  $\beta_1$  which is estimated by 2.03438

#### Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 168.60000 2.65702 63.45 < 2e-16 ***
time 2.03437 0.09039 22.51 2.16e-12 ***
```

Residual standard error: 3.234 on 14 degrees of freedom

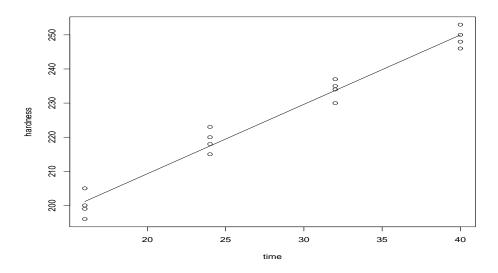


Figure 1: Plot of data and fitted line.

1b; problem 1.26 a.) In SAS you can get the residuals using the p option in proc reg; in R using the residuals function. Easy to see that they add to 0. Whenever we fit a regression model with an overall intercept the residuals will add to 0. This is not the case if we fit with no intercept.

```
hardness time fits resids
1 199 16 201.150 -2.150
2 205 16 201.150 3.850
```

```
3
        196
               16 201.150 -5.150
4
        200
               16 201.150 -1.150
5
        218
               24 217.425
                           0.575
6
        220
               24 217.425
                           2.575
7
        215
               24 217.425 -2.425
8
        223
               24 217.425
                           5.575
9
        237
               32 233.700
                           3.300
10
        234
               32 233.700
                           0.300
11
        235
               32 233.700
                           1.300
12
        230
               32 233.700 -3.700
13
               40 249.975 0.025
        250
14
        248
               40 249.975 -1.975
15
        253
               40 249.975 3.025
16
        246
               40 249.975 -3.975
```

b) The estimate of  $\sigma^2$  is  $\hat{\sigma}^2 = MSE = 10.45893$  (10.5 in the R output).

The estimate of  $\sigma$  is the square root of this so  $\hat{\sigma} = 3.23403$  (also root MSE on SAS output, Residual standard error in R output)).

1c) The (estimated) standard error for  $b_0$  is 2.657 and for  $b_1$  is .09039. The CI for  $b_0$  is computed using  $168.6 \pm t(.975, 14)2.657$  and for  $b_1$  using  $2.03435 \pm t(.975, 14).09039$ , where t(.975, 14) = 2.145. Using the confints in R or clb in SAS yields 95% confidence intervals of

```
(Intercept),\beta_0: [162.9013, 174.29875]
time, \beta_1: [1.8405, 2.22825]
```

## 1d, problems 2.7 a and b:

a) Asking for a 99% confidence interval for  $\beta_1$ , which is (1.7653, 2.30346). The interval can also be computed directly using  $b_1 \pm t(.995, 14)s\{b_1\}$  of obtained in either R (via confint with level = .99) or SAS (using clb with alpha = .01).

Using R

b) Testing  $H_0: \beta_1 = 2$  versus  $H_A: \beta_1 \neq 2$ .

The t-statistic is  $t^* = (2.03438 - 2)/0.09039 = .38035$ . With t(.975, 14) = 2.145, you do not reject  $H_0$  since .38 < 2.145. The P-value is the sum of the area to the right of .38 and the left of -.38 under the t distribution with 14 degrees of freedom. This actually equals .7094. Just using the t-tables, you can see that the area to the right of .38 is somewhere between .3 and .4, so from the tables you know the p-value is between .6 and .7. Note that we also accept accept  $H_0$  since the P-value is > .01.

- You can also test this directly using the 99% CI for  $\beta_1$  and reject  $H_0$  if 2 is not in the interval. Since 2 is in the interval we do not reject  $H_0$ . This is equivalent to doing the t-test.
- 1d: problem 2.16 You can get a) and b) in SAS directly from the output using the clm and cli option if you include a new case in the data with Y missing (.) and X = 30. In R and for the other parts, the intervals can be computed in various ways using either SAS or R as a calculator (as demonstrated for the Kishi example). See code and output at end of solution that corresponds to this.
- a)  $[227.4569, 231.8056] = 229.6313 \pm (2.264)0.8285$ , where  $s\{\hat{\mu}(30)\} = 0.8285$  is the standard error associated with the estimated mean and t(.99, 14) = 2.624.

Note that  $s^2\{\hat{\mu}(30)\} = 0.8285 = 7.06 + 30^2(.0082) + 2*30*(-.2288)$ , where  $s^2\{b_0\} = 7.06$ ,  $s^2\{b_1\} = .0082$  and  $s\{b_0, b_1\} = -.2288$  from the variance-covariance matrix of the coefficients

- b)  $(220.8695, 238.3930) = 229.6313 \pm (2.624)(10.45893 + 0.8285^2)^{1/2}$ .
- c)  $229.6313 \pm (2.264)((10.45893/10) + 0.8285^2)^{1/2} = [226.2, 233.1].$
- d) The interval in c) is smaller since you are trying to predict the mean of 10 values which has less variability  $(\sigma^2/10)$  than one value.
- e)  $229.6313 \pm (2*5.24)^{1/2}*0.8285 = [226.95, 232.32]$ , where F(.99, 2, 14) = 5.24 (obtained exactly using SAS or R you could approximate using the entries for F(.975, 2, 12), F(.975, 2, 15), F(.99, 2, 12) and F(.99, 2, 15) from the F table.)

1f.

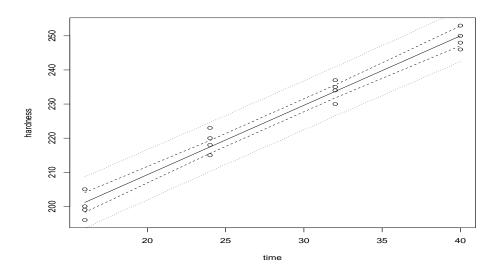


Figure 2: Individual CIs for means and Prediction intervals

# 1g.1g: Problem 4.9

a) and b)  $X_j = 20, 30$  and 40.

For Bonferroni, use  $b_0+b_1X_j\pm t(1-.10/6,14)s\{\hat{\mu}\{X_j\}\}$ , with t(1-.10/6,14)=2.360 The Working-Hotelling/Scheffe intervals use the same form but with the t value replaced by  $(2F(.90,2,14))^{1/2}=2.3352$ . The Working-Hotelling intervals will be smaller and more efficient since 2.335<2.360, but the difference is minor.

Scheffe intervals			Bonferro	ni intervals	estimate	SE
X=20	206.755	211.821	206.728	211.847	209.288	1.08473
X=30	227.697	231.566	227.676	231.586	229.631	0.82847
X = 40	246.816	253.134	246.783	253.168	249.975	1.35289

c). Now doing prediction intervals for two future values at  $X_1 = 30$  and  $X_2 = 40$ . Use  $b_0 + b_1 X_j \pm t(1 - .10/4, 14)s\{pred_j\}$ , with t(1 - .10/4, 14) = 2.145 and  $s^2\{pred_j\} = MSE + s^2\{\hat{\mu}(X_j)\}$ . The Scheffe intervals use the same form but with the t value replaces by  $(2F(.90, 2, 14))^{1/2} = 2.3352$ . The Bonferroni intervals should be used here since shorter.

```
Bonferroni Scheffe SEpredj
X=30 222.471, 236.792 221.836, 237.427 3.33846
X=40 242.456, 257.494 241.789, 258.161 3.50560
```

- 2. **Problem 2.** a) The estimate of  $\beta_0$  is .07471, of  $\beta_1$  is 2.10983, of  $\sigma^2$  is .06307, and of  $\sigma$  is .2511 =  $(.06307)^{1/2}$ .
  - b)  $2.10983 \pm 1.96(.01194) = [2.08643, 2.13323]$
  - c)  $H_0: \beta_1 = 0$ .  $t^* = 2.10983/.01194 = 176.67$ . The p-value for this is the area to the right of 176.67 plus the area to the left of 176.67 under the standard normal (use this since the degrees of freedom is 3164). Since the area to the right of 3.291 is equal to .0005 (see table) we know the P-value is less than  $2^*.0005 = .0001$ .

As a probability, the P-value is the probability of getting a value of the  $|t^*|$  greater than or equal to the value of the observed absolute value, under the null hypothesis. Here, it is the probability that  $|t^*| > 176.67$  where  $t^*$  is distributed t with 3164 degrees of freedom, which is for practical purposes the standard normal distribution.

d) 
$$\hat{Y}_h = 3.5917966$$
,  $s\{\hat{Y}_h\} = 0.0066024 = (s^2\{b_0\} + X_h^2 s^2\{b_1\} + 2X_h s\{b_0, b_1\})^{1/2}$ ,  $s\{pred\} = 0.2512242 = (MSE + s^1 2\{\hat{Y}_h\})^{1/2}$ .

Use z = 1.96 in getting the intervals.

Confidence interval for E(Y) at X = 1.667 is [3.5788559, 3.6047373]

Prediction interval for Y at X = 1.667, is [3.0993972, 4.084196].

3. Problem 3. Below are results for designs 1 and 2. Here I've given parts of the SAS output. Similar results apply for R. The first part of each program gives you the true standard errors of the estimated coefficients. The expected values of  $b_0$ ,  $b_1$  and MSE are known to be exactly  $\beta_0$  and  $\beta_1$  and  $\sigma^2$  (0, 1 and .0225) respectively.

To choose between designs we can do that in terms of the true standard errors of the estimated coefficients go. The second design is better because the standard errors are smaller. We know the estimators are unbiased for both unbiased so we can just compare variance or standard errors to make that decision.

- The second part of the output gives summary statistics over the thousands of simulated values. The fact that the means of the three variables are not exactly the true parameters is because we are running a limited number of simulations. The histograms represent the sampling distribution of these estimators (this is not the exact sampling distribution because of a limited number of simulations but it will be close).
- If you thought of which design is better for estimating  $\sigma^2$  you'd have to rely on the simulation results since I didn't tell you the variance of MSE. From the simulation results it looks like design 2 is a little better. In fact the two designs are equivalent from this perspective. It can be shown that the standard deviation of MSE is  $(2\sigma^4/(n-2))^{1/2}$  (= .01006 here), which only depends on the design through n. All designs with the same n are equally good for estimating  $\sigma^2$ . This may surprise you.

```
Homework 2, number 3 Design 1

coefficients: beta0 = 0 beta1 = 1

sigma2 0.0225 sigma = 0.15

number of simulations 1000

n = 12 xvalues = 7.67

6.31

6.14

7.07
```

6.39 5.95 6.53 6.55 5.34 5.74 4.94 7.07

## using normal errors

## THEORETICAL VARIANCES AND STANDARD DEVIATION

variance of b0 = 0.1413792 sd of b0 = 0.3760042 variance of b1 = 0.0035056 sd of b1 = 0.0592078

### The MEANS Procedure

Variable	N	Mean	Std Dev	Minimum	Maximum
B0	1000	0.0228333	0.3734848	-1.2547479	1.3339013
B1	1000	0.9966975	0.0588648	0.7961849	1.1947983
MSE	1000	0.0226470	0.0102871	0.0033132	0.0705808

And here are results from design 2.

Homework 2, number 3 Design 2

coefficients: beta0 = 0 beta1 = 1

5.3

5.3

1

5.3 5.3

6

6 7

7

7 7

## using normal errors

## THEORETICAL VARIANCES AND STANDARD DEVIATION

variance of b0 = 0.1181024 sd of b0 = 0.3436603 variance of b1 = 0.0030981 sd of b1 = 0.0556606

## The MEANS Procedure

Variable	N	Mean	Std Dev	Minimum	Maximum
В0	1000	0.0081894	0.3404172	-1.0216963	1.1265641
B1	1000	0.9987300	0.0549423	0.8055471	1.1717619
MSE	1000	0.0225139	0.0098290	0.0037846	0.0690602