

# Stat 525 Regression Analysis

## Lecture 1 : Linear Regression with One Predictor Variable

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# Outline

- Relations between Two Variables
- Regression Models and Their Uses
- Simple Linear Regression Model with Distribution of Error Terms Unspecified
- Overview of Steps in Regression Analysis
- Point estimation of  $E(Y) = \beta_0 + \beta_1 X$
- Point estimation of  $\sigma^2(Y) = \sigma^2$
- Simple Linear Regression Model with Normal Distribution Error Terms

# 1.1 Variable Types and Relations between Two Variables

## Dependent vs. Independent

- Independent variable( $X$ ) : predictor, explanatory variable
  - manipulated or changed by the experimenter
  - influences the dependent variable
- Dependent variable( $Y$ ) : response variable, outcome variable
  - observed result of the independent variable being manipulated
  - we want to predict

(e.g.) A call center where the number of customers serviced per hour, depends on the number of agents, and average service time per customer.

## Quantitative vs. Qualitative

- Quantitative variable
  - naturally measured as a number for which meaningful arithmetic operations make sense.
  - discrete variable and continuous variable
- Qualitative variable : categorical Variable
  - have no numerical meaning and take a value that is one of several possible categories

If  $X$  is an independent and quantitative variable and  $Y$  is a dependent and quantitative variable,

- Functional Relation :  $Y = f(X)$
- Statistical Relation :  $Y = f(X) + \epsilon$  where  $\epsilon$  is an (random) error term
  - variation in  $Y$  that is not associated with  $X$  and that is considered to be of a random nature
  - all data points do not fall directly on the line of relationship

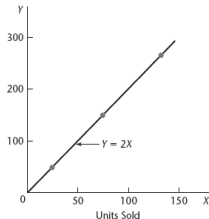


Figure 1.1(KNN)

Y(dollar sales of a product)

X(# of units sold)

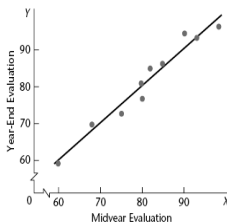


Figure 1.2(KNN)

Y(Year-end evaluation)

X(midyear evaluation)

FIGURE 1.3 Curvilinear Statistical Relation between Age and Steroid Level in Healthy Fem

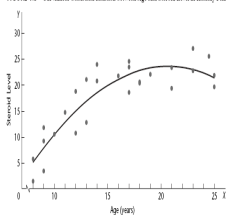


Figure 1.3(KNN)

Y(level of a steroid in plasma)

X(age)

## 1.2 Regression Models and Their Uses

### 1) Purpose of regression models

- determine the magnitude of the (typically imperfect) relationship between  $Y$  and a set of  $X$ s
- predict  $Y$  from a set of  $X$ s

### 2) Basic concepts

- A tendency of  $Y$  to vary with  $X$  in a systematic fashion
- A scattering of points around the curve of statistical relationship
  - Probability distribution of  $Y$  for each level of  $X$  :  $f(Y | X = x)$
  - **Regression function of  $Y$  on  $X$** ,  $E(Y | X) \equiv \int y f(Y | X) dy$  : the means of these probability distributions of  $Y$  vary in some systematic fashion with  $X$  and it is a function of  $X$
  - $Y = f(X) + \epsilon = E(Y | X) + \epsilon$

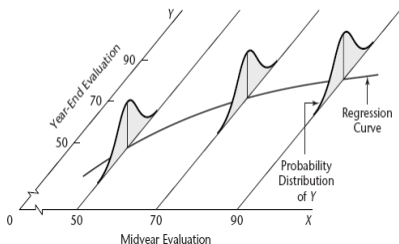


Figure 1.4(KNN)

### 3) Construction of Regression Models

- Selection of a set of “good”  $X$ s
- Functional form of regression relation
- Scope of regression model
- Regression and Causality

### 4) Data for regression analysis

- Observational data from nonexperimental studies that do not control  $X$ s of interest
  - no adequate information about cause-and-effect relationships
- Experimental data from experimental studies that control  $X$ s of interest through randomization
  - stronger information about cause-and-effect relationships
  - randomization balancing out the effects of other predictors that might affect  $Y$

## 1.3 Simple Linear Regression Model with Distribution of Error Terms Unspecified

$$Y_i = E(Y_i | X_i) + \epsilon_i = \beta_0 + \beta_1 X_i + \epsilon_i \quad i = 1, \dots, n \quad (1)$$

### ● Assumptions

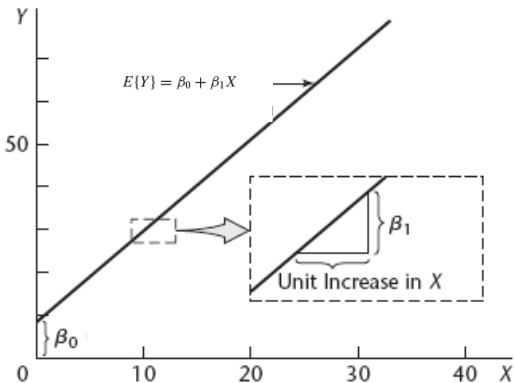
- $Y_i$  is the  $i$ -th value of the response variable
- $X_i$  is the  $i$ -th known value of the predictor variable (constant)
- $\beta_0$  and  $\beta_1$  are parameters (unknown constant) (regression coefficients)
- $\epsilon_i$  is an uncorrelated random error term with  $E(\epsilon_i) = 0$ ,  $\sigma^2(\epsilon_i) = \sigma^2$  and  $\sigma(\epsilon_i, \epsilon_j) = 0$

So,  $E(Y_i) =$  ,  $\sigma^2(Y_i) =$  ,  $\sigma(Y_i, Y_j) =$

- simple : there is only one  $X$  (multiple : # of  $X$  in the model  $> 1$ )
- linear in the parameters
- $\beta_0$ ,  $\beta_1$  and  $\sigma^2$  are the unknown parameters.

● Meaning of regression coefficients,  $\beta_0$  and  $\beta_1$

- $\beta_1$  = the slope (the change in the mean of the probability distribution of  $Y$  per unit increase in  $X$ )
- $\beta_0$  = the intercept (the mean of the probability distribution of  $Y$  at  $X = 0$ )





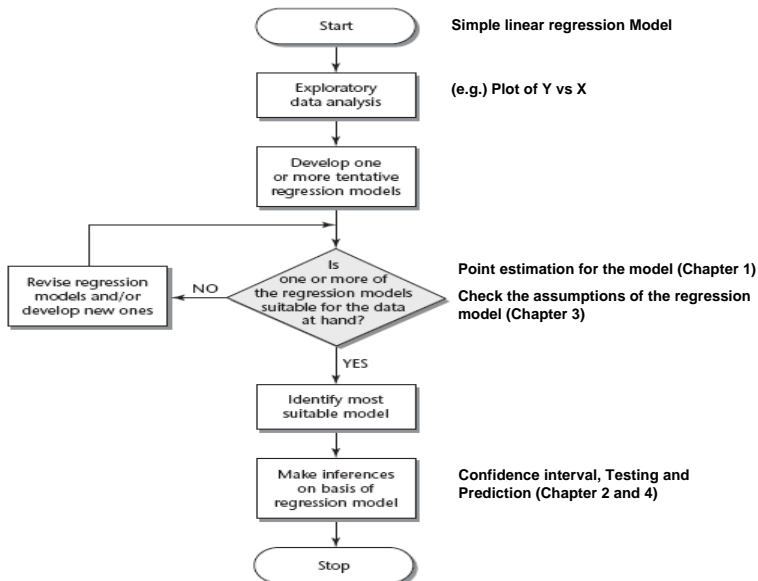
## • Features

- $Y_i$  is a random variable (why?)
  - mean response (regression function),  $E(Y_i) = \beta_0 + \beta_1 X_i$
  - $\sigma^2(Y_i) = \sigma^2$  : each probability distribution of  $Y$  has the same variance  $\sigma^2$
  - $\sigma(Y_i, Y_j)$  :  $Y_i$  and  $Y_j$  are uncorrelated
- $Y_i$  comes from probability distributions whose means are  $\beta_0 + \beta_1 X_i$  and whose variances are  $\sigma^2$ , the same for all levels of  $X$ . In addition,  $Y_i$  and  $Y_j$  are uncorrelated.

## • Alternative versions of $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$

- $Y_i = \beta_0 X_0 + \beta_1 X_i$  where  $\beta_0 \equiv 1$
- $Y_i = \beta_0^* + \beta_1(X_i - \bar{X})$  where  $\beta_0^* = \beta_0 + \beta_1 \bar{X}$

## 1.5 Overview of Steps in Regression Analysis



## 1.6 Estimation of $\beta_0$ , $\beta_1$ and $\sigma^2$

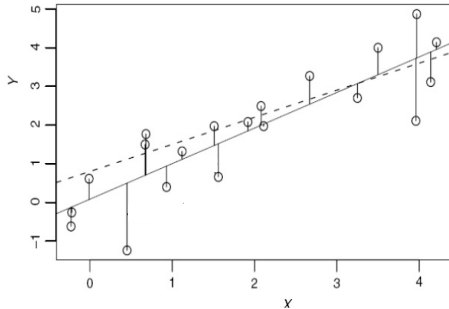
For the observations  $(X_1, Y_1), \dots, (X_i, Y_i), \dots, (X_n, Y_n)$ ,

- Use the **method of least squares** to obtain estimators of  $\beta_0$  and  $\beta_1$

**[Idea]** the estimators of  $\beta_0$  and  $\beta_1$  are those values  $b_0$  and  $b_1$ , respectively, minimizing  $Q$

$$Q = \sum_{i=1}^n \epsilon_i^2 = \sum_{i=1}^n (Y_i - E(Y_i))^2 = \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2$$

where  $Y_i - \beta_0 - \beta_1 X_i$  is the deviation of  $Y_i$  from its expected value



- Least Squares estimators for  $\beta_0$  and  $\beta_1$  are

$$b_1 = \frac{\sum_i (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_i (X_i - \bar{X})^2} = \frac{\sum_i (X_i - \bar{X})Y_i}{\sum_i (X_i - \bar{X})^2}, \quad b_0 = \bar{Y} - b_1\bar{X}$$

- How? solve  $\frac{\partial Q}{\partial \beta_0} = 0$  and  $\frac{\partial Q}{\partial \beta_1} = 0$  simultaneously

- Meaning of  $b_1$  and  $b_0$

(Study example) Suppose one is interested in the relationship between the number of hours ( $X$ ) given for study and score on a test ( $Y$ ). Given 20 observations  $(X_i, Y_i)$ , a simple linear regression was applied, and  $\beta_1$  and  $\beta_0$  using the method of least squares were calculated :  $b_1 = 3.5$  and  $b_0 = 15.05$

- students score \_\_\_ on average when they did not study
- adding an additional hour to your study time will result in an average score of \_\_\_ point higher

- Properties of  $b_0$  and  $b_1$

- $b_0$  and  $b_1$  are BLUE(Best Linear Unbiased Estimator)

- Point estimation of  $E(Y) = \beta_0 + \beta_1 X$

- Given  $b_0$  and  $b_1$ , the estimated regression function at  $X$  is

$$\hat{Y} = b_0 + b_1 X \quad (2)$$

so,  $\hat{Y}_i = b_0 + b_1 X_i$  where  $i = 1, \dots, n$  (called as the  $i$ -th fitted value)

(e.g.) In our (Study example),  $\hat{Y} = 15.05 + 3.5X$ . For a student studying 4 hours, the expected score on the exam is .

- Residuals,  $e_i = Y_i - \hat{Y}_i = Y_i - b_0 - b_1 X_i$

- vertical deviation of  $Y_i$  from the corresponding fitted value  $\hat{Y}_i$

- difference between  $\epsilon_i = Y_i - E(Y_i)$  and  $e_i = Y_i - \hat{Y}_i$

- very very useful for studying an estimated regression model is suitable for the  $n$  observations  $(X_i, Y_i)$  (Chapter 3)

- Properties of  $e_i$  and  $\hat{Y}_i$ 
  - $\sum_i e_i = 0$  and  $\sum_i e_i^2$  is a minimum
  - mean of  $\hat{Y}_i = \bar{Y}$ , i.e.,  $\frac{1}{n} \sum_i \hat{Y}_i = \frac{1}{n} \sum_i Y_i$
  - $\sum_i X_i e_i = 0$  and  $\sum_i \hat{Y}_i e_i = 0$
  - the regression line goes through the point  $(\bar{X}, \bar{Y})$
- Point Estimation of  $\sigma^2(Y) = \sigma^2(\epsilon) = \sigma^2$ 
  - $Y_i$  from different probability distributions with different means depending on  $X_i$
  - deviation of  $Y_i$  from  $\hat{Y}_i$ :  $e_i = Y_i - \hat{Y}_i$
  - $s^2 = MSE = \frac{SSE}{n-2} = \frac{\sum_{i=1}^n e_i^2}{n-2} = \frac{\sum_{i=1}^n (Y_i - \hat{Y}_i)^2}{n-2}$  where  $MSE$  is residual mean square and  $SSE$  is residual sum of squares : an estimator for  $\sigma^2$
  - $E(s^2) = \sigma^2$
  - $s = \sqrt{s^2}$  for the standard deviation  $\sigma = \sqrt{\sigma^2}$

# 1.8 Simple Linear Regression Model with Normal Distribution

## Error Terms

- Method of least squares
  - only know  $E(\epsilon_i) = 0$  and  $\sigma^2(\epsilon_i) = \sigma^2$  (the distribution of the  $\epsilon_i$  is unspecified)
  - $b_1$  and  $b_0$  are BLUE for  $\beta_0$  and  $\beta_1$  in Eq. (1), and  $s^2 = MSE = \frac{SSE}{n-2} = \frac{\sum_{i=1}^n e_i^2}{n-2}$  is an unbiased estimator for  $\sigma^2$
- One more assumption about the distribution of the  $\epsilon_i$  in  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$ 
  - need for interval estimators and hypothesis testing
  - $\epsilon_i \sim N(0, \sigma^2)$  and  $\sigma(\epsilon_i, \epsilon_j) = 0$  (uncorrelatedness implies independence between  $\epsilon_i$  and  $\epsilon_j$ ).

Then,  $Y_i \sim N(\beta_0 + \beta_1 X_i, \sigma^2)$ ,

and  $\sigma(Y_i, Y_j) = 0$

(so,  $Y_i$  and  $Y_j$  are independent).

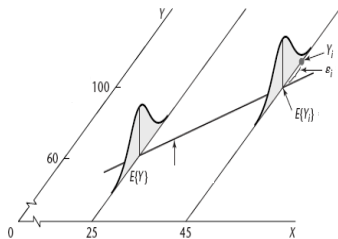


Figure 1.6 (KNN)

- Estimation of  $\beta_0$ ,  $\beta_1$  and  $\sigma^2$  in  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$  when  $\epsilon_i \sim N(0, \sigma^2)$  and  $\sigma(\epsilon_i, \epsilon_j) = 0$

· use Method of Maximum Likelihood

**[Idea]** construct the likelihood function of  $\beta_0$ ,  $\beta_1$  and  $\sigma^2$ ,  $L(\beta_0, \beta_1, \sigma^2)$ , and find values of the parameters maximizing the log of  $L(\beta_0, \beta_1, \sigma^2)$ ,  $\ell(\beta_0, \beta_1, \sigma^2)$

: Since  $Y_i \sim N(\beta_0 + \beta_1 X_i, \sigma^2)$ ,  $f(Y_i; \beta_0, \beta_1, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[ -\frac{1}{2} \left( \frac{Y_i - \beta_0 - \beta_1 X_i}{\sigma} \right)^2 \right]$ .

Then the (log) likelihood function is

$$L(\beta_0, \beta_1, \sigma^2) = \prod_{i=1}^n f(Y_i; \beta_0, \beta_1, \sigma^2) = \left( \frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left[ -\frac{1}{2\sigma^2} \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2 \right]$$

$$\ell(\beta_0, \beta_1, \sigma^2) = \log L(\beta_0, \beta_1, \sigma^2) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2$$

Then solve  $\frac{\partial \ell(\beta_0, \beta_1, \sigma^2)}{\partial \beta_0} = 0$ ,  $\frac{\partial \ell(\beta_0, \beta_1, \sigma^2)}{\partial \beta_1} = 0$  and  $\frac{\partial \ell(\beta_0, \beta_1, \sigma^2)}{\partial \sigma^2} = 0$  simultaneously.



- Estimators for  $\beta_0$ ,  $\beta_1$  and  $\sigma^2$  and their properties

Parameter	Method of Least Squares	Method of Maximum Likelihood
$\beta_1$	$b_1 = \frac{\sum_i (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_i (X_i - \bar{X})^2}$	$\hat{\beta}_1 = \frac{\sum_i (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_i (X_i - \bar{X})^2}$
$\beta_0$	$b_0 = \bar{Y} - b_1 \bar{X}$	$\hat{\beta}_0 = \bar{Y} - b_1 \bar{X}$
$\sigma^2$	$s^2 = \frac{\sum_{i=1}^n (Y_i - \hat{Y}_i)^2}{n-2}$	$\hat{\sigma}^2 = \frac{\sum_{i=1}^n (Y_i - \hat{Y}_i)^2}{n}$

- $\hat{\beta}_1$  and  $\hat{\beta}_0$  are unbiased, sufficient and consistent
- $\hat{\beta}_1$  and  $\hat{\beta}_0$  have minimum variance among all unbiased (linear or otherwise) estimators
- $s^2$  is unbiased, but  $\hat{\sigma}^2$  is biased for finite  $n$ .

## Calculate the regression coefficients and MSE with R

```
>x=c(-1,0,-2,-3)           #to introduce a variable
>y=c(-5,-4,2,-7)
>b_1=cov(x,y)/var(x)       #to calculate b1
>b_1                        #to call the value of b1
[1] 0.2
> b_0=mean(y)-b_1*mean(x)  #to calculate b0
> b_0                       #to call the value of b0
[1] -3.2
>yhat=b_0+b_1*x            #to introduce the regression model
>yhat                       #to call the value of yhat
[1] -3.4 -3.2 -3.6 -3.8
> n=length(y)              #to introduce the number of observations
> n                          #to call the value of n
[1] 4
> MSE=sum(y-yhat)^2/(n-2)  #to introduce MSE=S^2 as a variable
> MSE                       #to call the value of MSE
[1] 9.860761e-32
```