

UNIVERSITY OF MASSACHUSETTS  
 Department of Mathematics and Statistics  
 ADVANCED EXAM - LINEAR MODELS  
 January 15, 2010

Work all problems. 70 points are needed to pass.

1. (35 points) Suppose that we observe statistically independent observations with the same variance  $\sigma^2$ ,  $Y_{ij}$  for  $i = 1, 2$  and  $j = 1, 2$ , where  $E(Y_{ij})$  is modeled as follows:

$$E(Y_{11}) = \mu, \quad E(Y_{12}) = \mu + \gamma, \quad E(Y_{21}) = \mu + \alpha, \quad E(Y_{22}) = \mu + \alpha + \gamma,$$

- (a) With  $\boldsymbol{\beta} = (\mu, \alpha, \gamma)'$ , present this model in a linear regression form  $\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$ , and specify  $\mathbf{y}$ ,  $X$ , and  $\boldsymbol{\varepsilon}$ .
- (b) What is the rank of  $X$ ?, and what will be the d.f. of SSE?
- (c) Are  $\mu$ ,  $\alpha$ , and  $\gamma$  estimable? Give a reason.
- (d) Find a least squares (LS) estimate vector  $\hat{\boldsymbol{\beta}} = (\hat{\mu}, \hat{\alpha}, \hat{\gamma})'$  for  $\boldsymbol{\beta}$ , and write each of  $\hat{\mu}$ ,  $\hat{\alpha}$  and  $\hat{\gamma}$  explicitly as a linear combination of  $Y_{ij}$ 's. You may use

$$\begin{pmatrix} 4 & 2 & 2 \\ 2 & 2 & 1 \\ 2 & 1 & 2 \end{pmatrix}^{-1} = \begin{pmatrix} 3/4 & -1/2 & -1/2 \\ -1/2 & 1 & 0 \\ -1/2 & 0 & 1 \end{pmatrix}$$

- (e) Are those LS estimators  $\hat{\mu}$ ,  $\hat{\alpha}$  and  $\hat{\gamma}$  unique LS estimators for  $\mu$ ,  $\alpha$  and  $\gamma$  respectively? Further, are they the best linear unbiased estimators (BLUE) for  $\mu$ ,  $\alpha$  and  $\gamma$  respectively? Based on what theorem you make such conclusions?
- (f) Find the BLUE for  $2\mu + \gamma$  with explanations.
- (g) If  $Y_{11}$  is missing (not available), show that each of  $\mu$ ,  $\alpha$  and  $\gamma$  is still estimable.
2. (15 points)

- (a) Let  $\mathbf{Y}$  be a  $N_p(\boldsymbol{\mu}, \Sigma)$  random vector, where  $\Sigma$  is positive definite, and  $A$  be a  $p \times p$  symmetric matrix with rank  $k$  ( $k \leq p$ ). State a complete theorem concerning a necessary and sufficient condition for  $\mathbf{Y}'A\mathbf{Y}$  to be a chi-square distribution.
- (b) Let  $U_1, \dots, U_n$  be i.i.d. random vectors, where  $U_i = (x_i, y_i)'$  has a bivariate normal distribution with mean  $\boldsymbol{\mu} = (\mu_1, \mu_2)'$ ,  $\text{Var}(x_i) = \text{var}(y_i) = 1$ , and  $\text{Cov}(x_i, y_i) = \rho$  with  $\rho \neq -1$  or  $1$ , for  $i = 1, \dots, n$ . Define

$$q = n(\bar{x}^2 - 2\rho\bar{x}\bar{y} + \bar{y}^2)/(1 - \rho^2)$$

where  $\bar{x} = \mathbf{1}'\mathbf{x}/n$  and  $\bar{y} = \mathbf{1}'\mathbf{y}/n$  with  $\mathbf{x} = (x_1, \dots, x_n)'$  and  $\mathbf{y} = (y_1, \dots, y_n)'$ . Show that  $q$  has a chi-square distribution, and specify its degrees of freedom and non-centrality parameter. [Hint: Consider a quadratic form in

$$\mathbf{z} = \begin{pmatrix} \mathbf{x} \\ \mathbf{y} \end{pmatrix}. ]$$

3. (10 points) Consider the simple regression model:

$$Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i, \quad i = 1, \dots, n$$

where  $\varepsilon_i$  are independent with  $E(\varepsilon_i) = 0$ , and  $\text{Var}(\varepsilon_i) = \sigma^2 x_i^2$  (non-uniform variances) with  $\sigma^2 > 0$ . Find the BLUE for  $\beta_0$  and  $\beta_1$ , and an unbiased estimator for  $\sigma^2$ . (Note: You may give answers in matrix form, no need to be simplified.)

4. (20 points) Suppose we want to compare two simple linear models:

$$\begin{aligned} \text{Model I: } Y_{1j} &= \alpha_1 + \beta_1 x_{1j} + \varepsilon_{1j}, \quad j = 1, 2, 3 \\ \text{Model II: } Y_{2j} &= \alpha_2 + \beta_2 x_{2j} + \varepsilon_{2j}, \quad j = 1, 2, 3, 4 \end{aligned}$$

where all  $\varepsilon_{ij}$  are i.i.d.  $N(0, \sigma^2)$ .

- As a full model, write the two models in one form  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon}$  and define  $\mathbf{Y}$ ,  $X$ ,  $\boldsymbol{\beta}$  and  $\boldsymbol{\varepsilon}$ . Write down a LS  $\hat{\boldsymbol{\beta}}$  and SSE using these notations.
  - Suppose we want to test the null hypothesis  $H_0 : \alpha_1 = \alpha_2$  and  $\beta_1 = \beta_2$  against the alternative  $H_1 : \alpha_1 \neq \alpha_2$  or  $\beta_1 \neq \beta_2$ . Explain why  $H_0$  is a testable linear hypothesis if all matrices need to be inverted are non-singular.
  - Write down the reduced model (under  $H_0$ ) in the form  $\mathbf{Y} = Z\boldsymbol{\beta}_0 + \boldsymbol{\varepsilon}_0$ , and define  $\mathbf{Y}$ ,  $Z$ ,  $\boldsymbol{\beta}_0$  and  $\boldsymbol{\varepsilon}_0$ . Then write down a LS  $\hat{\boldsymbol{\beta}}_0$  and SSE<sub>0</sub> using these notations.
  - Write down the test statistic  $F$  and specify the criterion for rejecting  $H_0$  at 5% level.
5. (10 points) Let  $E_1, \dots, E_k$  be  $k$  events, then we have the Bonferroni inequality:

$$\Pr \left[ \bigcap_{i=1}^k E_i \right] \geq 1 - \sum_{i=1}^k \Pr[E_i^c].$$

We want to apply this inequality to construct simultaneous confidence intervals for  $k$  contrasts of means. Consider independent  $Y_{ij}$  for  $i = 1, 2, 3$  and  $j = 1, 2, \dots, n_i$ , where  $Y_{ij} \sim N(\mu_i, \sigma^2)$ . Describe how to construct simultaneous confidence intervals for  $\mu_2 - \mu_1$  and  $\frac{1}{2}(\mu_2 + \mu_3) - \mu_1$  with confidence at least 0.95. (Note: you need to write down explicit formulae for your notations.)

6. (10 points) Suppose we consider two competing models:  $Y = X\beta + \epsilon$  (reduced model) and  $Y = X\beta + Z\gamma + \eta$  (full model), where the error vectors ( $\epsilon$  and  $\eta$  respectively) follow a  $N(0, \sigma^2 \mathbf{I})$  distribution. Discuss how the estimated residuals from the reduced model can be used (through relevant plots) to assess model inadequacy. That is, what can the relevant residual plots reveal when the reduced model is correct and what can they reveal when the full model is correct. In particular discuss what happens when the reduced model is a simple linear regression and the full model is a quadratic regression of  $Y$  on  $X$ .